# Hypothesis Tests and Confidence Intervals in Multiple Regressions

SW Chapter 7

Overview of tests

Extended example: Angrist, Lang, and Oreopoulos (2009) Testing one restriction, multiple coefficients Testing multiple restrictions

Testing overall significance of a regression

- ▶ Construct and interpret tests of joint hypotheses
- ▶ Construct and test hypothesis test involving one restriction and multiple coefficients

Overview of tests

- 1. Hypothesis tests with one restriction, one coefficient
	- $▶$  Example:  $H_0$  :  $\beta_i = \beta_i$  vs.  $H_a$  :  $\beta_i \neq \beta_i$  o
- 2. Hypothesis tests with one restriction, multiple coefficients
	- $▶$  General: *H*<sub>0</sub> : *β*<sub>*i*</sub> = *β*<sub>*m*</sub>
	- $▶$  Example:  $H_0$  :  $\beta_1 = 0$
- 3. Hypothesis tests involving a multiple tests at once (joint hypothesis tests)
	- ▶ General: *H*<sup>0</sup> : *β<sup>j</sup>* = *βj,*0*, β<sup>m</sup>* = *βm,*0*, ...*
	- ▶ Example:  $H_0$ :  $β_1 = β_2 = β_3 = 0$
	- ▶ Special case test of *all* regressors

## 1. One restriction, one coefficient,  $\beta_i = \beta_{i,0}$

- 1. Select your significance level  $(\alpha = 0.01)$
- 2. State your null hypothesis:

$$
H_0: \beta_j = \beta_{j,0}
$$

$$
H_a: \beta_j \neq \beta_{j,0}
$$

3. Compute the t-statistic:

$$
t = \frac{\widehat{\beta}_j - \beta_{j,0}}{SE(\widehat{\beta}_j)}
$$

4. Compare the t-statistic to your critical value (2.58) and reject null if  $|t| > c$ 5. (Optional) Construct your confidence interval:

$$
(\widehat{\beta}_j-2.58SE(\widehat{\beta}_j),\widehat{\beta}_j+2.58SE(\widehat{\beta}_j)
$$

### 2. One restriction, two coefficients, *β<sup>j</sup>* = *β<sup>m</sup>*

- 1. Select your significance level  $(\alpha = 0.01)$
- 2. State your null hypothesis:  $H_0$ :  $\beta_1 = \beta_2$  vs  $H_a$ :  $\beta_1 \neq \beta_2$
- 3. Transform your regression:

$$
y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + u_i
$$
  
\n
$$
y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{1i} - \beta_2 x_{1i} + \beta_2 x_{2i} + u_i
$$
  
\n
$$
y_i = \beta_0 + (\beta_1 - \beta_2) x_{1i} + \beta_2 (x_{1i} + x_{2i}) + u_i
$$
  
\n
$$
y_i = \beta_0 + \gamma_1 x_{1i} + \beta_2 w_i + u_i
$$

and instead test  $H_0$ :  $\gamma_1 = 0$  vs.  $H_1$ :  $\gamma_1 \neq 0$ 

4. Repeat remaining steps

There are lots of variants, but this is the one we will compute by hand

- 1. Select your significance level ( $\alpha = 0.01$ )
- 2. State your null hypothesis:  $H_0$ :  $\beta_1 = \beta_2 = \beta_3$  vs any part not true
- 3. Estimate model with variables you are testing (unrestricted) and without (restricted)
- 4. Calculate *F*-statistic

$$
F = \frac{(SSR_r - SSR_{ur})/q}{SSR_{ur}/(n - k_{ur} - 1)} = \frac{(R_{ur}^2 - R_r^2)/q}{(1 - R_{ur}^2)/(n - k_{ur} - 1)} \sim F_{q, n - k_{ur} - 1}
$$

5. Compare the F-statistic to your critical value from a *Fq,n−kur−*<sup>1</sup> critical value and reject null if  $F > c$  (usually, this will be from  $F_{a,\infty}$  distribution)

Extended example: Angrist, Lang, and Oreopoulos (2009)

### Angrist, Lang, and Oreopoulos (2009)

#### Incentives And Services For College Achievement: Evidence From A Randomized Trial

By JOSHUA ANGRIST, DANIEL LANG, AND PHILIP OREOPOULOS\*

This paper reports on an experimental evaluation of strategies designed to improve academic performance among college freshmen. One treatment group was offered academic support services. Another was offered financial incentives for good grades. A third group combined both interventions. Service use was highest for women and for subjects in the combined group. The combined treatment also raised the grades and improved the academic standing of women. These differentials persisted through the end of second year, though incentives were given in the first year only. This suggests study skills among some treated women increased. In contrast, the program had no effect on men. (*JEL* 121, 128)

#### We're going to think about predictors of year 1 college GPA.

### A satellite campus of a "large, Canadian university"



with controls for gender, mother tongue, and HS quartile (no top quartile!)

# Avoid dummy variable trap!





English mother tongue excluded:

$$
G\widehat{PA} \widehat{year1} = 1.484 - 0.130
$$
 *female* – 0.466 *mt<sub>french</sub>* – 0.076 *mt<sub>other</sub>* + 0.374 *hs<sub>q2</sub>* + 0.886 *hs<sub>q3</sub>*

"Other" mother tongue excluded:

$$
G\widehat{PA} \widehat{year1} = 1.409 - 0.130
$$
 *female* – 0.076 *mt<sub>english</sub>* – 0.390 *mt<sub>french</sub>* + 0.374*hs<sub>q2</sub>* + 0.886*hs<sub>q3</sub>*

## Choice of excluded group only affects interpretation

For simplicity, assume *female* = 0,  $hs_{q2} = 0$ , and  $hs_{q3} = 0$ 

English mother tongue excluded:

$$
G\widehat{PA} \text{year1} = 1.484 - 0.130 \text{female} - 0.466 m t_{\text{french}} - 0.0755 m t_{\text{other}} + 0.374 h s_{q2} + 0.886 h s_{q3}
$$

$$
E[GPAyear1|mt_{english} = 1] = 1.484
$$
  
\n
$$
E[GPAyear1|mt_{french} = 1] = 1.484 - 0.466 = 1.018
$$
  
\n
$$
E[GPAyear1|mt_{other} = 1] = 1.4843 - 0.0755 = 1.409
$$

## Choice of excluded group only affects interpretation

For simplicity, assume *female*  $= 0$ ,  $hs_{q2} = 0$ , and  $hs_{q3} = 0$ 

"Other" mother tongue excluded:

*GPAyear* \ <sup>1</sup> <sup>=</sup> <sup>1</sup>*.*<sup>409</sup> *<sup>−</sup>* <sup>0</sup>*.*130*female* <sup>+</sup> <sup>0</sup>*.*0755*mtenglish <sup>−</sup>* <sup>0</sup>*.*390*mtfrench*  $+ 0.374$ *hs*<sub>*q*2</sub> + 0.886*hs*<sub>*q*3</sub>

$$
E[GPAyear1|mt_{english} = 1] = 1.4087 + 0.0755 = 1.484
$$
  
\n
$$
E[GPAyear1|mt_{french} = 1] = 1.4087 - 0.3903 = 1.018
$$
  
\n
$$
E[GPAyear1|mt_{other} = 1] = 1.4087 = 1.4087 = 1.409
$$

#### Questions we could want to answer



. regress GPA vearl female mt french mt other hs g2 hs g3

- ▶ Do French speakers have same GPA as English speakers?
- Do non-English/French speakers have same GPA as English speakers?
- Do French speakers have same GPA as non-French/English speakers?
- ▶ Are there differences in GPA by mother tongue?

# Do French speakers have same GPA as non-French/English speakers? Population model:

 $G$ *FAyear*<sup>1</sup>*i* =  $\beta_0 + \beta_1 m t_{\text{french}} + \beta_2 m t_{\text{other}} + \beta_3 h s_{q2} + \beta_4 h s_{q3} + u$ 

Test whether  $H_0$ :  $\beta_1 = \beta_2$  vs  $H_a$ :  $\beta_1 \neq \beta_2$ Equivalent to asking whether  $\beta_1 - \beta_2 = 0$ 

#### Do French speakers have same GPA as non-French/English speakers?

```
. test mt_french = mt_other
(1) mt french - mt other = 0F( 1, 1368) = 1.59Prob > F = 0.2073
```
#### Are there differences in GPA by mother tongue?

Population model:

 $G$ *FAyear*<sup>1</sup>*i* =  $\beta_0 + \beta_1 m t_{\text{french}} + \beta_2 m t_{\text{other}} + \beta_3 h s_{q2} + \beta_4 h s_{q3} + u$ 

Test  $H_0: \beta_1 = \beta_2 = 0$  vs.  $H_a$  any one equality is not true (either  $\beta_1 \neq 0$  or  $\beta_2 \neq 0$ ) Equivalent to asking whether the coefficients are jointly significant

- ▶ Exclusion restriction Test of whether certain covariates can be excluded from population model
- ▶ Restricted model: The model with *fewer* covariates
	- $\blacktriangleright$  "Restricted" because we are imposing that coefficients of missing covariates are zero
- ▶ Unrestricted model The model with *more* covariates
	- ▶ "Unrestricted" because coefficients on those covariates can be zero or anything else

Idea: How would the model fit be if these variables were dropped from the regression? We can consider how much variation is unexplained using *SSR*

Source	SS	df	MS	Number of obs	$=$	1.374
				F(5, 1368)	$=$	57.06
Model	186.639466	5	37.3278932	$Prob$ > $F$	$=$	0.0000
Residual	894.929455	1,368	.654188198	R-squared	$=$	0.1726
				Adj R-squared	$=$	0.1695
Total	1081.56892	1.373	.787741385	Root MSE	$=$	.80882
GPA_year1	Coef.	Std. Err.	t	P >  t		[95% Conf. Interval]
female	$-.1296654$	.0442666	$-2.93$	0.003 $-.2165032$		$-.0428276$
mt french	$-.4658419$	.3079241	$-1.51$	0.131 $-1.069896$		.1382127
mt other	$-.0755215$	.04796	$-1.57$	0.116 $-.1696047$		.0185616
$hs$ q2	.3743258	.0545463	6.86	0.000 .2673224		.4813292
hs a3	.8857912	.0532672	16.63	0.000 .781297		.9902854
$_{\rm -cons}$	1.484259	.0467597	31.74	0.000 1.392531		1.575987

. regress GPA yearl female mt french mt other hs q2 hs q3

The sum of squared residuals *necessarily* increases, but is that increase statistically significant?



$$
F = \frac{(SSR_r - SSR_{ur})/q}{SSR_{ur}/(n - k_{ur} - 1)} \sim F_{q,n-k-1}
$$

- $\blacktriangleright$  *q* is the number of restrictions you're examining.
- ▶ *n − kur −* 1 is degrees of freedom in unrestricted model

The relative increase of *SSR* going from restricted to unrestricted follows a *F*-distribution if  $H_0$  is true.

- ▶ A *F*-distributed variable takes on positive values, reflects fact that *SS* can only increase as we add more variables (move from restricted to unrestricted)
- ▶ Choose critical value so that the null hypothesis is falsely rejected in *α*% of cases (type I error)
- $▶$  For now, let  $\alpha = 0.1$

$$
F = \frac{(897.91 - 894.93)/2}{894.93/(1374 - 5 - 1)} = 2.28
$$
  

$$
F \sim F_{2,1368} \rightarrow c_{0.1} = 2.30
$$
  

$$
P(F > 2.28) = 0.1032
$$

We cannot reject null at conventional levels.

The relative increase of *SSR* going from restricted to unrestricted follows a *F*-distribution if  $H_0$  is true.

. regress GPA\_year1 female mt\_french mt\_other hs\_q2 hs\_q3

Source	SS	df	MS		Number of obs	$=$	1.374
					F(5, 1368)	$=$	57.06
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Residual	894.929455	1,368	.654188198		R-squared	$=$	0.1726
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Total	1081.56892	1.373	.787741385	Root MSE		$=$	.80882
GPA_year1	Coef.	Std. Err.	t	P >  t			[95% Conf. Interval]
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mt french	$-.4658419$	.3079241	$-1.51$	0.131	$-1.069896$		.1382127
mt_other	$-.0755215$	.04796	$-1.57$	0.116	$-.1696047$		.0185616
hs q2	.3743258	.0545463	6.86	0.000	.2673224		.4813292
$hs_q3$	.8857912	.0532672	16.63	0.000	.781297		.9902854
cons	1.484259	.0467597	31.74	0.000	1.392531		1.575987

```
. test mt_french = mt_other=0
```

```
(1) mt_french - mt_other = 0
(2) mt_french = 0
     F(2, 1368) = 2.28Prob > F = 0.1032
```

```
. testparm material measurement materials and material measurement of the state of the state
```
. regress GPA\_year1 female mt\_french mt\_other hs\_q2 hs\_q3,robust Linear regression





```
. test mt french = mt other=0
(1) mt french - mt other = 0(2) mt french = 0
     F(2, 1368) = 2.86Prob > F = 0.0573. testparm mt_french mt_other
(1) mt_french = 0
(2) mt other = 0
     F(2, 1368) = 2.86Prob > F = 0.0573
```
#### Test of overall significance of a regression

- ▶ Can we exclude all the regressors?
- $\triangleright$  General population model (unrestricted):

$$
y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \ldots + \beta_k x_{ik} + u_i
$$

$$
\blacktriangleright \text{ Test } H_0: \beta_1 = \beta_2 = ... \beta_k = 0
$$

▶ Restricted model:

$$
y_i = \beta_0 + u_i
$$

$$
F = \frac{(SSR_r - SSR_{ur})/q}{SSR_{ur}/(n-k-1)} \sim F_{q,n-k-1} = \frac{R^2/k}{(1 - R^2)/(n-k-1)} \sim F_{k,n-k-1}
$$

The test of overall significance is reported in most regression packages. Usually, null hypothesis is overwhelmingly rejected.

#### Test of overall significance of a regression

. regress GPA\_year1 female mt\_french mt\_other hs\_q2 hs\_q3





- . testparm \*
- $(1)$  female = 0  $(2)$  mt french = 0  $(3)$  mt other = 0  $(4)$  hs\_q2 = 0  $(5)$  hs  $q3 = 0$  $F( 5, 1368) = 57.06$  $Prob > F = 0.0000$
- ▶ T-test: test of just one regressor, one-sided or two-sided
- ▶ F-test: test of just one regressor or multiple regressors, two-sided only

When working with just one regressor,  $F=t^2$ , so results are functionally equivalent!

# Conclusion

### Overview of tests

Extended example: Angrist, Lang, and Oreopoulos (2009)

- Testing one restriction, multiple coefficients
- Testing multiple restrictions
- Testing overall significance of a regression